Continuous Branching process with Immigration conditioned to stay positive

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jointly with Víctor Rivero (CIMAT)

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- BGWI (Bienaymé-Galton-Watson process with immigration) is described by two generating functions f and g, where f corresponds to regeneration and g corresponds to immigration.
- Under specific scalings (and possibly considering a sequence of processes) BGWIs are known to converge to Continuous Branching processes with immigration (CBI), see [Kawazu and Watanabe 1971].

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- Under specific scalings (and possibly considering a sequence of processes) BGWIs are known to converge to Continuous Branching processes with immigration (CBI), see [Kawazu and Watanabe 1971].

• In [Caballero, Pérez Garmendia, and Uribe Bravo 2013] it was shown that a CBI process Z is a solution to a time-change equation:

$$Z_t = X\left(\int_0^t Z_s ds\right) + Y_t,\tag{1}$$

where \boldsymbol{Y} is a subordinator and \boldsymbol{X} is a spectrally positive Levy process with finite mean.

Consider the generating functions

 $f(s) := s + c(1-s)^{\alpha+1}L(1-s)$ and $g(s) := 1 - d(1-s)^{\alpha}G(1-s),$

where $c,\ d>0$ and $\alpha\in(0,1]$ and $L(x)\sim G(x)$ for $x\to 0$ and L is slowly varying at 0.

Note that the BGWI(f, g) is critical: f'(1) = 1, and

 $f''(1) < \infty$ iff $g'(1) < \infty$ iff $\alpha = 1$ and $G(0) < \infty$.

Proposition

Let η_{α} be a r.v. with a tail $\mathbb{P}(\eta_{\alpha} > x) \sim dx^{-\alpha}G(x)$ as $x \to \infty$. Then

 $g \sim Pois(\eta_{\alpha})$ and $f \sim Be(1/Pois(\eta_{\alpha})) \cdot (Pois(\eta_{\alpha}) + 1)$.

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 $g \sim Pois(\eta_{\alpha})$ and $f \sim Be(1/Pois(\eta_{\alpha})) \cdot (Pois(\eta_{\alpha}) + 1).$

Denote by \mathfrak{Z} the BGWI(f,g). Define b_n by $b_n^{\alpha} \sim nL(1/b_n)$ (as $n \to \infty$) and let

$$Z_t^{(n)} := \frac{1}{b_n} \mathfrak{z}(\lfloor nt \rfloor).$$
⁽²⁾

It holds that (in Skorokhod's space)

$$Z^{(n)} \xrightarrow{w} Z, \ n \to \infty,$$

where Z is an α -self-similar CBI: the regenerating and immigration mechanisms are given by

$$R(q) = -cq^{\alpha+1}$$
 and $F(q) = dq^{\alpha};$

or, equivalently, Z satisfies the time-change equation $Z_t = X\left(\int_0^t Z_s ds\right) + Y_t$ with Y being an α -stable subordinator and X being a $1 + \alpha$ -stable spectrally positive Levy process.

We assume that

$$\delta := \frac{d}{\alpha c} \in (0, 1). \tag{3}$$

This assumption guarantees that 0 is a recurrent point of Z, [Foucart and Uribe Bravo 2014].

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The scaling limit

Denote by \mathfrak{Z} the $\mathrm{BGWI}(f,g)$. Define b_n by $b_n^{\alpha} \sim nL(1/b_n)$ (as $n \to \infty$) and let

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- **9** Yaglom limit: conditioning on non-extinction up till time u > 0.
- Ourrett's extension to the meander.
- $\textbf{ S Letting } u \to \infty.$
- **4** Recognising the *h*-transform of the original process.

Yaglom limit

Let $d_0(w)$ denote the hitting time of 0 for w:

$$d_0(w) \coloneqq \inf\{s > 0 : w(s) = 0\}.$$

Consider

$$\lim_{n \to \infty} \mathbb{P}\left(\frac{1}{b_n} \mathfrak{z}(n) \in A \mid d_0(\mathfrak{z}) > n\right).$$

Theorem (Yaglom limit, [Mijatović, P., and Uribe Bravo 2025])
$$\lim_{n \to \infty} \mathbb{E}\left(\exp\left(-\lambda \frac{\tilde{z}(n)}{b_n}\right) \mid d_0(\tilde{z}) > n\right) = \frac{1}{(1 + \alpha c \lambda^{\alpha})}.$$

Compare to the similar limit for a BGW \tilde{z} , [Slack 1968]:

$$\lim_{n \to \infty} \mathbb{E}\left(\exp\left(-\lambda \frac{\tilde{\beta}(n)}{b_n}\right) \mid d_0(\tilde{\beta}) > n\right) = 1 - \left(\frac{\alpha c \lambda^{\alpha}}{1 + \alpha c \lambda^{\alpha}}\right)^{1/\alpha}$$

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Consider a Markov chain $(v(k))_{k\geq 0}$ on \mathbb{R} and assume that

- Scaling limit: for $(b_n)_n$ and $z_n := v(0)/b_n \to z$, it holds $Z_n^{z_n} := v(n \cdot)/b_n \xrightarrow{w} Z^z$.
- **2** Enough mass beyond 0: for t, z > 0, $\mathbb{P}(\inf_{s < t} Z^{z}(t) > 0) > 0$.
- Convergence of hitting times of 0: $t_n \to t > 0$ and $z_n \to z > 0$, then $\mathbb{P}(d_0(Z_n^{z_n}) > t_n) \to \mathbb{P}(d_0(Z^z) > t)$.
- Regularity of 0: $t_n \to t > 0$ and $z_n \to 0$, then $\mathbb{P}(d_0(Z_n^{z_n}) > t_n) \to 0$.

Theorem (**[Durrett 1978])

Assume 1-4 hold. Fix u > 0 and let $Z_n^{(+,u)} \coloneqq (Z_n^0 \mid d_0(Z_n) > u)$. Then

- $\exists \beta > 0$ such that $\mathbb{P}(d_0(Z_n) > u) \sim n^{-\beta} L^*(n)$ for a slowly varying L^* ;
- (Z_n^(+,u)(s))_{s∈[0,u]} converges weakly (in D[0,∞)) to an inhomogeneous Markov process Z^(+,u).

Define κ_u as $Z_n^{(+,u)}(u) \xrightarrow{w} \kappa_u$. By the self-similarlity of the limit, this will satisfy $\kappa_u := u^{1/\alpha} \kappa_1$. The transition density of $Z^{(+,u)}$ is given by:

for t > 0

$$\mathbb{P}_0(Z^{(+,u)}(t) \in dy) = (u/t)^{\beta} \mathbb{P}((t/u)^{1/\alpha} \kappa_u \in dy) \mathbb{P}_z(d_0(Z) > u - t);$$

for s < t < u and z > 0

$$\mathbb{P}(Z^{(+,u)}(t) \in dy \mid Z^{(+,u)}(s) = z) \\ = \mathbb{P}_z \begin{pmatrix} Z(t-s) \in dy, \\ d_0(Z) > t-s \end{pmatrix} \frac{\mathbb{P}_y(d_0(Z) > u-t)}{\mathbb{P}_z(d_0(Z) > u-s)}.$$

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Consider $Z_n^{(+,u)} := (\mathfrak{z}(\lfloor n \cdot \rfloor)/b_n \mid d_0(\mathfrak{z}) > nu)$. There exists an inhomogeneous Markov process $Z^{(+,u)}$ such that

$$Z_n^{(+,u)} \xrightarrow{w} Z^{(+,u)}, \text{ as } n \to \infty.$$

Constant β is given by [Mijatović, P., and Uribe Bravo 2025] and it is equal to $\beta := 1 - \delta$.

From Yaglom limit we find that κ_u is defined as

$$\mathbb{E}e^{\lambda\kappa_u} = \frac{1}{(1 + \alpha c u \lambda^\alpha)}.$$

Define a random variable κ as $\mathbb{E}e^{-\lambda\kappa} = (1 + \lambda^{\alpha})^{-1}$. Then

 $\mathbb{P}_0(Z^{(+,u)}(t) \in dy) = (u/t)^{1-\delta} \ \mathbb{P}((\alpha ct)^{1/\alpha} \kappa \in dy) \ \mathbb{P}_y(d_0(Z) > u-t).$

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Letting $u \to \infty$: part I

For
$$s < t < u$$
 and $z > 0$
$$\lim_{u \to \infty} \mathbb{P}(Z^{(+,u)}(t) \in dy \mid Z^{(+,u)}(s) = z)$$
$$= \mathbb{P}_z \begin{pmatrix} Z(t-s) \in dy, \\ d_0(Z) > t-s \end{pmatrix} \lim_{u \to \infty} \frac{\mathbb{P}_y(d_0(Z) > u-t)}{\mathbb{P}_z(d_0(Z) > u-s)}.$$

Proposition (by-product of [Mijatović, P., and Uribe Bravo 2025])

Let $A \sim Arcsin(\delta) \sim Beta(\delta, 1 - \delta)$ and independent $E \sim Exp(1)$. Then

$$d_0(Z^z) \stackrel{d}{=} \frac{z^{\alpha}}{\alpha c} \frac{E^{-\alpha}}{A}.$$

For $h(z) = z^{\alpha(1-\delta)}$, we have

$$\mathbb{P}_z(d_0(Z) > u) \sim \frac{\Gamma(1-\alpha)}{\Gamma(\delta)\Gamma(2-\delta)} \frac{h(z)}{(\alpha c u)^{1-\delta}}, \quad \text{as } u \to \infty$$

Hence

$$\mathbb{P}(Z^{(+,u)}(t) \in dy \mid Z^{(+,u)}(s) = z) \to \mathbb{P}_z(Z(t-s) \in dy, d_0(Z) > t-s) h(y)/h(z).$$

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For $h(z) = z^{\alpha(1-\delta)}$, we have

$$\mathbb{P}_z(d_0(Z)>u)\sim \frac{\Gamma(1-\alpha)}{\Gamma(\delta)\Gamma(2-\delta)}\frac{h(z)}{(\alpha cu)^{1-\delta}}, \quad \text{as } u\to\infty.$$

Hence

$$\mathbb{P}(Z^{(+,u)}(t) \in dy \mid Z^{(+,u)}(s) = z) \to \mathbb{P}_z(Z(t-s) \in dy, d_0(Z) > t-s) h(y)/h(z).$$

There exists a (time-homogeneous) Markov process $Z^{(+)}$, strictly positive on $(0,\infty)$, such that the finite-dimensional distributions of $Z^{(+,u)}$ converge to those of $Z^{(+)}$ as $u \to \infty$:

$$Z^{(+,u)} \stackrel{fdd}{\to} Z^{(+)}, \quad \text{as } u \to \infty.$$
(4)

The process $Z^{(+)}$ is an *h*-transform of Z with $h(z) := z^{\alpha(1-\delta)}$ and it has the following transition densities: for y > 0, z > 0, 0 < t < u

$$\mathbb{P}_{z}(Z^{(+)}(t) \in dy) = \mathbb{P}_{z}(Z(t) \in dy, d_{0}(Z) > t) \frac{h(y)}{h(z)};$$
(5)

for y > 0 and z > 0 and 0 < s < t

$$\mathbb{P}(Z^{(+)}(t) \in dy \,|\, Z^{(+)}(s) \in dz) = \mathbb{P}_z(Z^{(+)}(t-s) \in dy);$$

for z = 0 and t > 0

$$\mathbb{P}_0(Z^{(+)}(t) \in dy) = \frac{\Gamma(1-\alpha)}{\Gamma(\delta)\Gamma(2-\delta)} \frac{h(y)}{(\alpha ct)^{1-\delta}} \mathbb{P}((\alpha ct)^{1/\alpha} \kappa \in dy).$$
(6)

<ロト < 団ト < 巨ト < 巨ト < 巨ト 三 の Q (~ 12/22 To see that the function h is a Doob h-transform we need to verify that it is excessive for the process killed at the hitting time of 0, i.e., $\mathbb{E}_{z} [h(Z(t)); d_{0}(Z) > t] \uparrow h(z)$ as $t \downarrow 0$, we note that for any z > 0

$$\mathbb{E}_{z}\left[\frac{h(Z(t))}{h(z)}; d_{0}(Z) > t\right] = \mathbb{P}_{z}^{(+)}\left[1\right] \le 1,$$
(7)

where $\mathbb{P}_z^{(+)}$ is the probability measure associated with $Z^{(+)}$ issued from z. As $t \downarrow 0$, by Lebesgue's Dominated convergence theorem, we see that $\mathbb{E}_z [h(Z(t)); d_0(Z) > t] \uparrow h(z)$ as required.

Consider the Lamperti transform of Z

$$Z(t)1_{\{d_0(Z)>t\}} = ze^{\xi(\tau(tz^{-\alpha}))}, \text{ where } \tau_t := \inf\{s>0: \int_0^s e^{\alpha\xi(u)}u > t\}.$$
(8)

Since Z hits zero continuously, $\lim_{t\to\infty} \xi_t = -\infty$.

Theorem

The Laplace transform of ξ is given by

$$\Psi(\lambda) := \log\left(\mathbb{E}e^{-\lambda\xi}\right) = \frac{\Gamma(\alpha+\lambda)}{\Gamma(\lambda)}\alpha c(1-\delta+\lambda/\alpha).$$

Moreover, we find that $\mathbb{E}e^{\alpha(1-\delta)\xi} = 1$, i.e., the Cramer's constant of ξ is $\alpha(1-\delta)$ and also $\left(e^{\alpha(1-\delta)\xi_t}\right)_{t\geq 0}$ is a martingale.

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Theorem (Volkonskii's theorem)

The generator of Z and the generator of ξ are related through

$$\mathcal{L}_Z f(z) = z^{-\alpha} \mathcal{L}_{e^{\xi}} f(z) = z^{-\alpha} \mathcal{L}_{\xi} \tilde{f}(\log(z)),$$

where $\tilde{f}(z) := f(e^z)$.

Being an h-transform implies that $h(Z(t))1_{\{d_0(Z)>t\}}$ is a supermartingale. It is actually a martingale.

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The process
$$(h(Z(t))1_{\{d_0(Z)>t\}})_t = (h(ze^{\xi(\tau(tz^{-\alpha}))}))_t$$
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• From [Bertoin and Doney 1994], as $x \to \infty$

$$\mathbb{P}_{z}(\max_{t \le d_{0}(Z)} Z(t) > x) = \mathbb{P}(z \max_{t \ge 0} e^{\xi(t)} > x) \sim \frac{C_{1}/C_{2}}{\alpha(1-\delta)} \frac{h(z)}{h(x)},$$

where $C_1 := -\log \mathbb{P}(H_1 < \infty)$ and $C_2 := \mathbb{E}(H_1 e^{\alpha(1-\delta)H_1}; H_1 < \infty)$ with H being the ascending ladder height process of ξ .

• As
$$x \to \infty$$

 $\mathbb{P}_0(Z^{(+)}(t) > x) \sim \frac{\sin(\pi\delta)}{\pi\delta(1-\delta)} (\alpha ct)^{\delta} x^{-\alpha\delta}$

<ロト 4 回 ト 4 臣 ト 4 臣 ト 臣 の Q () 16 / 22 Following [Noba 2023], we can define the $q\mbox{-scale}$ function of a spectrally positive process Z as

$$W_Z^{(q)}(x,y) = \begin{cases} 1/N^x \left[e^{-qT_y^-} \right], & y \le x, \\ 0, & y > x, \end{cases}$$

for x, y in $(0, \infty)$; where N^x is the excursion measure (from x) and T_y^- is the hitting time of the [0, y].

$$\mathbb{E}_{x}\left[e^{-qT_{a}^{-}};T_{a}^{-} < T_{b}^{+}\right] = \frac{W_{Z}^{(q)}(b,x)}{W_{Z}^{(q)}(b,a)}$$

Following [Noba 2023], we can define the $q\mbox{-scale}$ function of a spectrally positive process Z as

$$W_Z^{(q)}(x,y) = \begin{cases} 1/N^x \left[e^{-qT_y^-} \right], & y \le x, \\ 0, & y > x, \end{cases}$$

for x, y in $(0, \infty)$; where N^x is the excursion measure (from x) and T_y^- is the hitting time of the [0, y].

$$\mathbb{E}_{x}\left[e^{-qT_{a}^{-}};T_{a}^{-} < T_{b}^{+}\right] = \frac{W_{Z}^{(q)}(b,x)}{W_{Z}^{(q)}(b,a)}$$

By [Noba 2023], we see that the q-scale function $(W_Z^{(q)}(x,y))_x$ of Z, for $q \ge 0$, is the unique solution to the Volterra integral equation

$$f(x) = x^{\alpha - 1} W_{\xi}(\log(y/x)) + x^{\alpha - 1} q \int_{(x,y)} f(x') W_{\xi}(\log(x'/x)) dx'.$$

In particular, for the $0\mbox{-scale}$ function of Z is immediately given by

$$W_Z(x,y) = x^{\alpha-1} W_{\xi}(\log(y/x))$$

Proposition

The 0-scale function of ξ is given by

$$W_{\xi}(x) = \frac{e^{-\alpha(1-\delta)x}}{\Gamma(1+\alpha)c(1-\delta)} \int_{0}^{1-e^{-x}} u^{\alpha-1}(1-u)^{-1-\alpha(1-\delta)} u$$
$$= \frac{e^{-\alpha(1-\delta)x}(1-e^{-x})^{\alpha}}{\Gamma(1+\alpha)\alpha c(1-\delta)} {}_{2}F_{1}(\alpha, 1+\alpha(1-\delta); 1+\alpha; 1-e^{-x}).$$

<ロト < 部ト < 差ト < 差ト 差 の < で 18/22 By [Noba 2023], we see that the q-scale function $(W_Z^{(q)}(x,y))_x$ of Z, for $q \ge 0$, is the unique solution to the Volterra integral equation

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<ロト < 部ト < 言ト < 言ト 言 の Q (~ 18/22 $\boldsymbol{\xi}$ is a Lévy process hence the scale function satisfies

$$W_{\xi}^{(q)}(x-y) = W_{\xi}^{(q)}(x,y) \text{ and } \int_{0}^{\infty} e^{-\lambda x} W_{\xi}^{(q)}(x) dx = \frac{1}{\Psi(\lambda) - q}.$$

For q = 0, we compute

$$W_{\xi}(x) = \mathcal{L}^{-1}\left[\frac{1/\Gamma(1+\alpha)}{c(1-\delta)}\frac{\Gamma(\lambda)\Gamma(\alpha)}{\Gamma(\alpha+\lambda)}\frac{\alpha(1-\delta)}{\alpha(1-\delta)+\lambda}\right](x).$$

Finally, recall the Beta function

$$\int_0^\infty e^{-\lambda x} (1 - e^{-x})^{\alpha - 1} x = \frac{\Gamma(\lambda)\Gamma(\alpha)}{\Gamma(\lambda + \alpha)}.$$

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Thank you for your attention!



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