Coexisting of branching populations Branching and persistence 2025

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9th of April, Angers

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- Branching processes
- Random walks in cones

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- Main results

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Branching process in random environment

Random environment

We consider a random environment sequence Q = (Q(1), Q(2), ...) of random probability measures on \mathbb{N}_0 . We also consider it to be i.i.d. thus Q(1), Q(2), ... are independent copies of some random probability measure μ with generating function Fon \mathbb{N}_0 .



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Branching process

Branching processes in random environment could be thought of as developing in a scheme of two stages. First, the random environment Q = (Q(1), Q(2), ...) is established. Given its value q = (q(1), q(2), ...), the branching process \mathcal{Z} then evolves in the varying environment Q. Z(n) may again be realized as

$$Z(n) = \sum_{i=1}^{Z(n-1)} Y(i, n),$$

where given the environment q(n), Y(i, n) are random independent variables with distributions q(n).

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Random walks

Associated random walk

Let Q be an i.i.d. environment. We require F[0] < 1 and $0 < \mathbf{E}\mu < \infty$ a.s. Then we can define $X(n) := \log \mathbf{E}q(n)$ and

$$S(0) = 0, \quad S(n) = \sum_{i=1}^{n} X(n)$$

 ${S(n)}_{n=0}^{\infty}$ is called an associated random walk.

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 $\{S(n)\}_{n=0}^{\infty}$ is called an associated random walk.

Critical Branching process

We say that Branching process in random environment is critical if

$$\limsup_{n\to\infty} S(n) = \infty, \ \liminf_{n\to\infty} S(n) = -\infty$$

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Two-dimensional case 0000 000

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Two-dimensional case 0000 000

Survival probability

Survival probability in critical case

There exists a positive finite constant $\mathcal Y$ such that, as $n \to \infty$:

$$\mathbb{P}(Z(n) > 0) \sim \mathcal{Y} \cdot \mathbb{P}(\min\{S(0), \dots, S(n)\} \ge 0)$$



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Two-dimensional case

Survival probability

Survival probability in critical case

There exists a positive finite constant \mathcal{Y} such that, as $n \to \infty$:

$$\mathbb{P}ig(Z(n)>0ig)\sim \mathcal{Y}\cdot\mathbb{P}ig(\min\{S(0),\ldots S(n)\}\geqslant 0ig)$$

Stopping time

We define

$$\tau_x := \inf\{n \ge 1 : x + S(n) \le 0\} \ x \ge 0.$$

Then we can reformulate the result above as:

$$\mathbb{P}(Z(n) > 0) \sim \mathcal{Y} \cdot \mathbb{P}(\tau_0 > n).$$

Here we can consider stopping time τ_0 as a first time random walk $\{S(n)\}_{n=0}^{\infty}$ leaves the cone $\mathcal{K}^1 := \{y \in \mathbb{R} : y \ge 0\}$ if it starts at S(0) = 0.

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Harmonic functions

Renewal function

In the critical case we have important result. Let V(x) be a renewal function for random walk $\{S(n)\}$. Then one has $\mathbb{P}(\tau_x > n) \sim \frac{V(x)}{\sqrt{n}}$ (Doney '95). For the renewal function in one-dimensional case main asymptotics is known

$$V(x) \sim C \cdot x, \quad x \to \infty,$$

where C > 0 is some constant that does not depend on x and n.

Fwo-dimensional case

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Harmonic functions

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$$V(x)\sim C\cdot x,\quad x
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where C > 0 is some constant that does not depend on x and n.

Harmonic property

V(x) is strictly positive in \mathcal{K}^1 and moreover for arbitrary $x \in \mathcal{K}^1$ it is true that:

$$\mathbb{E}\big[V(x+X), \ x+X \in \mathcal{K}^1\big] - V(x) = 0,$$

where X is one step of random walk. In other words, V is a strictly positive eigenfunction for operator

$$f \rightarrow \mathbb{E}f - f$$
,

with eigenvalue 0.

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Multi-type Branching processes in random environment

One-dimensional Branching processes in joint environment

Here we consider several population with the following properties:

- particles from different populations can only produce particles from its own population;
- one-dimensional environments for different populations could be dependent.

For two-dimensional case we can represent $\{(Z_1(n), Z_2(n))\}_{n=0}^{\infty}$ in the following way:

$$(Z_1(n), Z_2(n)) = \sum_{i=1}^{Z_1(n-1)} (Y_1(i, n), 0) + \sum_{j=1}^{Z_2(n-1)} (0, Y_2(j, n)).$$

Critical case

We are interested in the case when both one-dimensional branching processes are critical.

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One-dimensional case

Two-dimensional case ○●○○ ○○○

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Branching processes

Main example

Consider two one-dimensional i.i.d. environments $(Q_1(1), Q_1(2), ...)$ and $(Q_2(1), Q_2(2), ...)$ with corresponding generating functions $(F_1^{(1)}, F_1^{(2)}, ...)$ and $(F_2^{(1)}, F_2^{(2)}, ...)$:

$$F_1^{(i)}(s_1) = \sum_{k=1}^{\infty} \mathbf{P}(Q_1(i) = k) s_1^k = p_1(i) + \sum_{k=1}^{\infty} (1 - p_1(i))^k p_1(i) s_1^k$$

$$F_2^{(j)}(s_2) = \sum_{k=1}^{\infty} \mathbf{P}(Q_2(j) = k) s_2^k = p_2(j) + \sum_{k=1}^{\infty} (1 - p_2(j))^k p_2(j) s_2^k$$

Then $\mathbf{E}Q_1(1) = rac{p_1}{1-\rho_1}$ and $\mathbf{E}Q_2(1) = rac{p_2}{1-\rho_2}$.

$$X_1(i) := \log \mathsf{E}Q_1(i), \ X_2(i) := \log \mathsf{E}Q_2(i), \ S_1(n) := \sum_{k=1}^n X_1(k) \quad S_2(n) := \sum_{k=1}^n X_2(k),$$

where we set

$$\operatorname{Cor}[X_1(1)X_2(1)] = \varrho, \quad \varrho \in (1,1).$$

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We also consider $\mathbb{E}[X_1(1)^2] = \mathbb{E}[X_2(1)^2] = 1$.

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Global goal

Survival probability and coexistence

We are considering which is called **coexistence** probability: $\mathbb{P}(Z_1(n) > 0, Z_2(n) > 0)$.



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Global goal

Survival probability and coexistence

We are considering which is called **coexistence** probability: $\mathbb{P}(Z_1(n) > 0, Z_2(n) > 0)$.

Main example

For one-dimensional case we can represent $\mathbb{P}(Z(n) > 0))$ in a more convenient form:

$$\mathbb{P}(Z(n) > 0)) = \left(\frac{1}{\mathsf{E}Q(1)\dots\mathsf{E}Q(n)} + \sum_{k=1}^{n} \frac{\phi(Q(1),\dots,Q(n))}{\mathsf{E}Q(1)\dots\mathsf{E}Q(n)}\right)^{-1}$$

Recalling definition of S(k) and using the expression above for the main example we can simplify:

$$\mathbb{P}(Z_1(n) > 0, Z_2(n) > 0) = \mathbb{E}\left[\left(\sum_{k=1}^n e^{-S_1(k)}\right)^{-1}\left(\sum_{k=1}^n e^{-S_2(k)}\right)^{-1}\right].$$

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Two-dimensional case 000● 000

Condition to stay in cone

Main idea

One can see that

$$\begin{split} \mathbb{E}\bigg[\big(\sum_{k=1}^{n} e^{-S_{1}(k)}\big)^{-1}\big(\sum_{k=1}^{n} e^{-S_{2}(k)}\big)^{-1}\bigg] \\ & \geq \mathbb{E}\bigg[\big(\sum_{k=1}^{n} e^{-S_{1}(k)}\big)^{-1}\big(\sum_{k=1}^{n} e^{-S_{2}(k)}\big)^{-1}, \ \min_{k \leq n} \{S_{1}(k), S_{2}(k)\} > -R\bigg] \\ & \times \mathbb{P}\big(\min_{k \leq n} \{S_{1}(k), S_{2}(k)\} > -R\big). \end{split}$$

Stopping time

For positive quadrant \mathcal{K}^2 let τ_x be the first time *n* such that x + S(n) leaves \mathcal{K}^2 . Then we can rewrite the expression above as

$$\mathbb{E}\left[\left(\sum_{k=1}^{n} e^{-S_1(k)}\right)^{-1}\left(\sum_{k=1}^{n} e^{-S_2(k)}\right)^{-1}, \ \tau_{(R,R)} > n\right] \cdot \mathbb{P}(\tau_{(R,R)} > n)$$

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Random walks in cones

Main asymptotic

Let \mathcal{K}^2 be a strictly positive quadrant in \mathbb{R}^2 . Let also $Cor(X_1(1), X_2(1)) = \varrho$. Then:

$$\mathbb{P}(au_x > n) \sim V(x) rac{1}{n^{ heta(arrho)/2}},$$

where $\varrho \in (-1, 1)$ and $\theta(\varrho)$ is some function from ϱ . (V. Wachtel, D. Denisov '15)



Two-dimensional case

Random walks in cones

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Harmonic property and upper bound

V(x) is strictly positive in \mathcal{K}^2 and for arbitrary $x \in \mathcal{K}^2$ one has

$$\mathbb{E}\big(V(x+X), \quad x+X \in \mathcal{K}^2\big) - V(x) = 0.$$

Moreover, for arbitrary $x \in \mathcal{K}^2$ it is true that

$$V(x) \leqslant C \cdot |x|^{\theta(\varrho) - 1} \cdot dist(x, \partial \mathcal{K}^2)$$

Main results



First result

Theorem 1

Assume that $|\varrho| < 1$. Set

$$\theta(\varrho) = \frac{\pi}{2 \arccos(-\varrho)}.$$

Assume also that $\mathbb{E}|X(n)|^{2\theta}$ and $\mathbb{E}|X(n)|^2 \log(1 + |X(n)|)$ are finite. Then for every starting point z for branching process as in the main example there exists a positive constant A = A(z) such that

$$\mathbb{P}_{z}(Z_{1}(n) > 0, Z_{2}(n) > 0) \sim An^{-\theta(\varrho)}$$
 as $n \to \infty$. (1)

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Second result

Theorem 2

On the set $E(n) := \{Z_1(n) > 0, Z_2(n) > 0\}$ we define

$$z^{(n)}(t)=\left(rac{\log Z_1(nt)}{\sqrt{n}},rac{\log Z_1(nt)}{\sqrt{n}}
ight),\quad t\in[0,1].$$

Suppose all the conditions of Theorem 1 hold. Then the sequence $z^{(n)}$ conditioned on E(n) converges in distribution on D[0,1] with uniform metric. The limiting process can be called Brownian meander in \mathcal{K}^2 .

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Main results



Idea of the proof of the Theorem 1

Doob h-transform

For arbitrary $v \in \mathcal{K}^2$ one can consider Doob h-transform:

$$\widehat{\mathbb{P}}(x+S(n)=v):=\frac{V(v)}{V(x)}\mathbb{P}(x+S(n)=v,\tau_x>n).$$

We want to estimate $\hat{\mathbb{P}}(x_1 + S_1(n) < h(n))$ for some function $h(\cdot)$.



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Estimation in crucial points

First, we are estimating small deviations of the form $\hat{\mathbb{P}}(x + S_2(n) < h(n), x + S_1(n) \leq \chi(n))$. Main estimates are obtained from small enough values of V(v).

Second, we are estimating large deviations of the form

 $\hat{\mathbb{P}}(x + S_2(n) < h(n), x + S_1(n) > \chi(n)$. Main estimates are obtained from Fuk-Nagaev inequalities.

This estimates guarantee us that

$$\mathbb{E}\left[(\sum_{k=1}^{n} e^{-S_1(k)})^{-1} (\sum_{k=1}^{n} e^{-S_2(k)})^{-1}, \ \tau_{(R,R)} > n \right].$$

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